Discussion: Methodologies for the new censuses

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Discussion

Two papers

- S. Falorsi: Census and social surveys integrated systems
 - \blacksquare Combining administrative and survey data \longrightarrow Permanent Census
 - Software package MIND
- M. Di Zio and D. Filipponi: *Multi-source statistics in the Italian permanent census*
 - \blacksquare Combining administrative and survey data \longrightarrow Imputation of ALE and OCC
 - Variance estimation

- Italian permanent census: combine administrative and survey data
- Reasons:
 - Significant reduction of the costs;
 - Reduction of the respondent's burden;
- ISTAT developed a new data methodological/statistical framework by integrating 3 components:
 - Integrated Register System (IRS): Integrates data from administrative sources and surveys at the individual level → the missing data are imputed
 - Permanent Population Census (PPC): produces set of estimates (via small area techniques) that cannot be obtained through administrative sources
 - Census and Social Surveys Integrated System (CSSIS): Other set of estimates (SAE based on unit level models)

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Discussion

- PPC: Goal is to produce a set of values (observed or predicted) at the individual level → micro-level approach
- Requirements for the set of estimates produced by the PPC:
 - Accuracy (Validity): this requires the specification of a model → Model diagnostics become central

$$\mathsf{MSE}(\widehat{\theta}) = \mathbb{V}(\widehat{\theta}) + \mathsf{Bias}^2(\widehat{\theta})$$

- Efficiency: small variance
- \blacksquare Consistency: internal and external consistency \longrightarrow Specify a set of calibration constraints

- In PPC/CSSIS, estimation procedures involve SAE methods;
- Impressive *R*-package Mind (Multivariate model-based INference for Domains)
- Can handle multiple survey variables but with a common/unique set of covariates (Limitation?)
- Allows for different correlation structures (including spatial correlations between levels of each random effect)
- Estimation of MSE

- Since with data integration methods, we make many assumptions, a lot of efforts should be placed on model diagnostics to detect departures from model assumptions
- For instance, the macro GEST developped at Statistics Canada includes a SAE component. GEST offers several useful diagnostics for the Fay-Herriot model (only the combined Fay-Herriot model can be validated):
 - Plot of residuals vs. set of predictors, predicted values, etc.; If the assumptions do not seem to be satisfied, then we may consider adding polynomial term of higher order or consider piecewise linear regression;
 - Plot of standardized square residuals vs. set of predictors, predicted values; If the assumptions do not seem to be satisfied, then try to determine the right amount of heteroscedasticity (may not be easy);
 - Normality of the standardized errors;
 - Before validating the Fay-Herriot model, important to validate the smoothing model (that was used to smooth the sampling variances)

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- Does MIND propose (or will propose) a set of diagnostics in the context of multivariate models? May be useful to users.
- Other important issue in SAE: Outliers
 - Outlier detection (influential domains or influential units);
 - Important issue in unit level models;
 - May be also an issue in area level models;
 - Vast literature on robust SAE for unit level models (e.g., Bertarelli et al., 2022, Favre-Martinoz, 2015, Dongmo Jiongo et al., 2013, Sinha and Rao, 2009) but not much has been done for area level models;
 - \blacksquare What about robust multivariate SAE models \longrightarrow More research needed?

- Mass imputation/propensity score weighting are common in the data integration context;
- The resulting estimators are vulnerable to model misspecification \longrightarrow nonparametric/machine learning methods may bring some robustness
- Often we make a MAR-type assumption. What if it's not satisfied?
 - We can have recourse to multiply robust procedures (e.g., Chen and Haziza, 2022): In the case of NMAR, these methods tend to lead to a better estimator (although inconsistent) than the one that would have been obtained under a single misspecified model;
 - Multiply robust propensity score weighting assuming NMAR (Kim and Cho, 2022) → the price to pay is to have an independent validation sample with the same measurements (y and x) → There is no free lunch!

- Goal: Mass impute the variable Attained Level of Education (ALE) and Occupational Status (OCC) in the Italian Base Register of Individuals
- Variables used to mass impute come from:
 - Ministry of Education Universities and Research (MIUR): administrative data;
 - 2011 Census Information;
 - Sample Survey (collected since 2018);
- Mass imputation is justified by the high amount of detailed information \longrightarrow Rich imputation model

- Different patterns of missing data involve different set of covariates
- Imputation procedure:
 - First, estimate P(ALE^t | x) using a log-linear model applied to the contingency table obtained by cross-classifying the variables ALE^t and x → P(ALE^t | x)
 - Randomly generate a ALE status with probability $\widehat{P}(ALE^t | x)$
 - If we use a saturated model, then equivalent to random hot-deck imputation within cells
 - \blacksquare If some cells are empty \longrightarrow use a subset of covariates selected through a cross-validation procedure

Patterns of missing data

X _{BRI}				Xmiur				Sample	Prediction	Group
G	E	Р	Ct	L(t)	ALE^{t-12}	F(t)	ALE ^{apr}	ALE_{S}^{t}	ALE^{t}	
										А
										22%
										В
										73%
										С
										5%

Variance estimation: Numerical results (Di Zio, Filippini, Toti, 2022)

ALE	Analytical	Margin of error
	\widehat{V}	$1.96\sqrt{\widehat{V}}$
Illiterate	$1.42 imes10^{-8}$	0.000233
Literate but no attainment	$6.62 imes10^{-8}$	0.000504
Primary education	$1.82 imes10^{-7}$	0.000836
Lower secondary	$3.78 imes10^{-7}$	0.001250
Upper secondary	$3.69 imes10^{-7}$	0.001190
BSc	$7.49 imes10^{-8}$	0.000536
MSc	$9.81 imes10^{-8}$	0.000613
PhD	$1.16 imes10^{-8}$	0.000211

Table 1: Variance estimates and associated margin of errors

- Extremely small margins of error!
- May be due to:
 - Very large sample sizes;
 - \blacksquare Very powerful covariates \longrightarrow Imputation model highly predictive
- Point estimator:

$$\widehat{\overline{Y}} = \frac{1}{N} \left\{ \sum_{i \in S} y_i + \sum_{i \in S^c} \widetilde{y}_i \right\} = \frac{n}{N} \overline{y}_s + \left(1 - \frac{n}{N} \right) \overline{\widetilde{y}} \approx \overline{\widetilde{y}}$$

With such small variances, a small bias may lead to invalid inferences
→ coverage probability of normal-based confidence intervals may be
much lower than 95% even if the bias is small → Importance of
validating the model as much as possible

- Imputation of OCC is more complex as all the data sources may suffer from measurement errors → Use of mixture Markov models
- Variance estimation for OCC:
 - Analytical approach too complicated;
 - \blacksquare Use of Multiple Imputation (MI) \longrightarrow MILC procedure (Boeschoten et al., 2020)
- Multiple Imputation variance estimator (Rubin, 1978):

$$\widehat{V}_{tot} = \overline{W}_M + \left(1 + \frac{1}{M}\right) B_M,$$

where

$$\overline{W}_M = \frac{1}{M} \sum_{m=1}^M W^{(m)}, \qquad B_M = \frac{1}{M-1} \sum_{m=1}^M \left(\widehat{\theta}_l^{(m)} - \widehat{\theta}_{l,M} \right)^2$$

• Validity of \hat{V}_{tot} relies of the fact that the imputation procedure is proper.

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